

Simulated Maximum  
Likelihood for  
Continuous-Discrete  
State Space Models  
using  
Langevin Importance  
Sampling.

Hermann Singer

State Space Models

Parameter  
Estimation

Langevin Sampling

Simulated Likelihood

Examples

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# Simulated Maximum Likelihood for Continuous-Discrete State Space Models using Langevin Importance Sampling.

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Estimation of Stochastic Differential Equations with Time  
Series, Panel and Spatial Data.

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# Nonlinear continuous/discrete state space model

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$$dY(t) = f(Y(t), t)dt + g(Y(t), t)dW(t)$$

$$Z_i = h(Y_i, t_i) + \epsilon_i$$

$$i = 0, \dots, T$$

- nonlinear **drift and diffusion** functions  $f, g$   
 $f = f(Y(t), t, x(t), \psi)$
- nonlinear diffusion  $g(Y)$ : **Itô calculus**
- **Spatial models:**  $Y_n(t) = Y(x_n, t), x_n \in \mathbb{R}^d$ :  
Random field  $Y(x, t, \omega)$

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# Linear stochastic differential equations (LSDE)

(exact ML, Singer; 1990, Kalman filter)

$$\begin{aligned} dY(t) &= AY(t)dt + GdW(t) \\ Y(t) &= e^{A(t-t_0)} Y(t_0) + \int_{t_0}^t e^{A(t-s)} GdW(s) \end{aligned}$$

- $W(t, \omega)$ : Wiener process:  
continuous time random walk (Brownian motion)
- Itô stochastic differential equations:  
 $dW/dt = \zeta(t)$  = Gaussian white noise
- $A$ : drift matrix
- $G$ : diffusion matrix

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# Wiener process and stock index

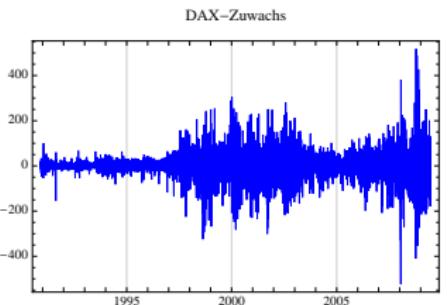
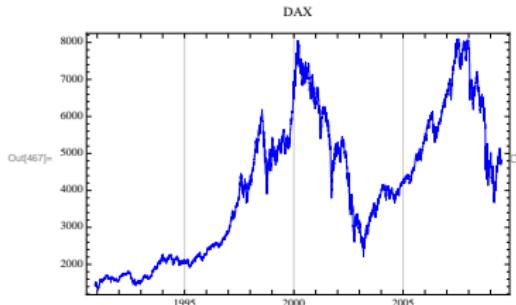


Figure : German stock index (DAX)

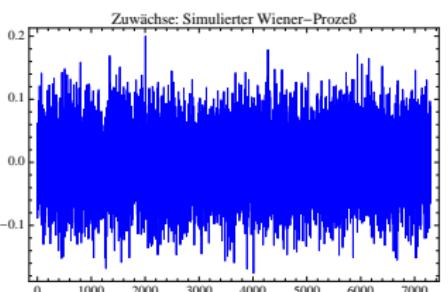
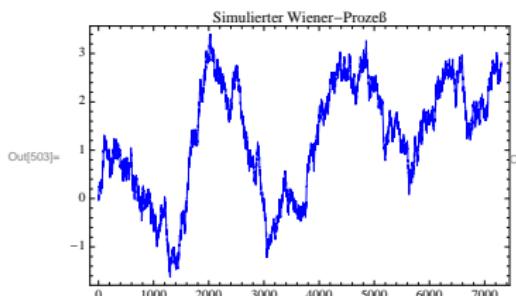


Figure : Simulated Wiener process (random walk)

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# Simulated Wiener processes

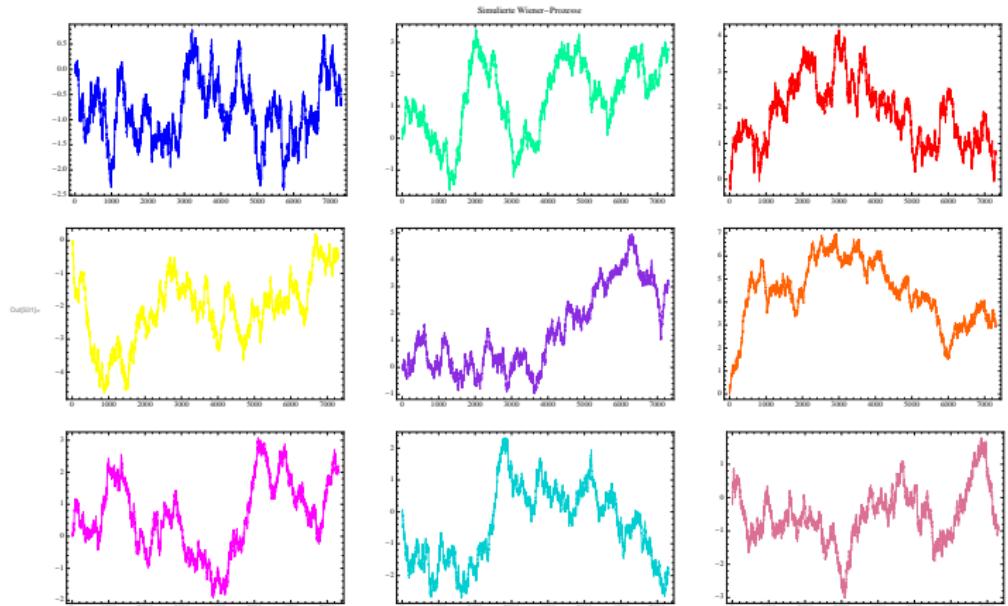


Figure : Simulated Wiener processes

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# Exact discrete model (EDM)

Bergstrom (1976, 1988)

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$$Y_{i+1} = e^{A(t_{i+1}-t_i)} Y_i + \int_{t_i}^{t_{i+1}} e^{A(t_{i+1}-s)} G dW(s)$$

restricted VAR(1) model

$$Y_{i+1} = \Phi(t_{i+1}, t_i) Y_i + u_i$$

- $\Phi$ : fundamental matrix of the system
- $Y_i := Y(t_i)$ : sampled measurements

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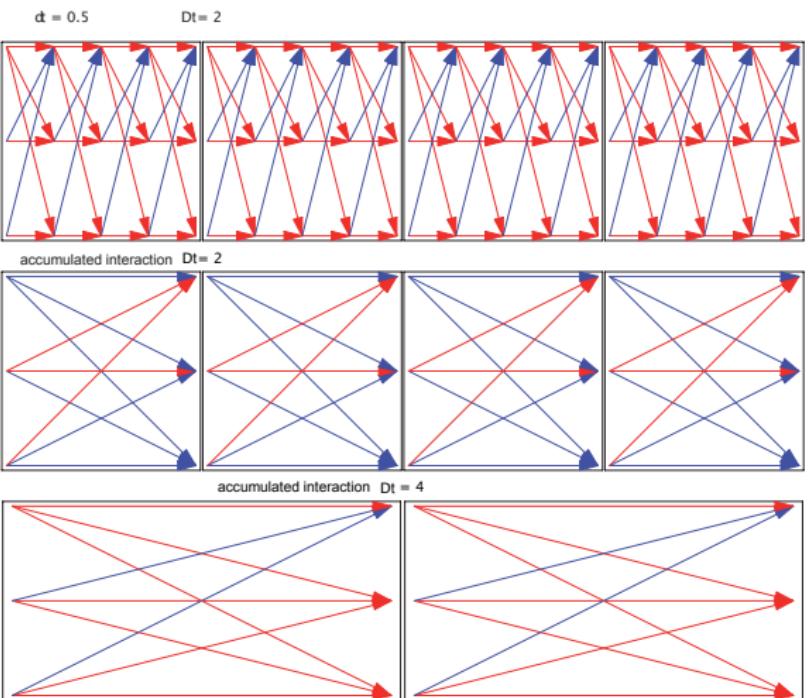


Figure : 3 variable model:

Product representation of matrix exponential within the measurement interval  $\Delta t = 2$ . Latent states  $\eta_j$ , discretization interval  $\delta t = 2/4 = 0.5$  (Singer; 2012)

# Maximum Likelihood Parameter Estimation

likelihood function of all observations

$$p(z_T, \dots, z_0) = \int p(z_T, \dots, z_0 | y_T, \dots, y_0) p(y_T, \dots, y_0) dy$$

- High dimensional integration over latent variables
- **smooth** dependence on parameter vector  $\psi$ :  
 $L(\psi) = p(z; \psi)$
- Problem:  $p(y_T, \dots, y_0)$  not known  
Sampling interval  $\Delta t_i$ ;  
Transition density  $p(y_{i+1}, \Delta t_i | y_i)$  difficult to compute
- Use additional latent variables  
 $y_T = \eta_J, \dots, \eta_0 = y_0, y(t_i) = y_i = \eta_j;$

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# Integration

latent states  $\eta_j$

$$\begin{aligned} p(z_T, \dots, z_0) &= \int p(z_T, \dots, z_0 | \eta_J, \dots, \eta_0) p(\eta_J, \dots, \eta_0) d\eta \\ &= E[p(z_T, \dots, z_0 | \eta_J, \dots, \eta_0)] \end{aligned}$$

- $\eta_{j_i} = y(t_i) = y_i$ ,  $t_i$  = measurement times
- even higher (infinite) dimensional integration over latent variables:  
**Markov Chain Monte Carlo: simulate  $\eta$**
- Euler density (discretization interval  $\delta t$ )

$$p(\eta_{j+1}, \delta t | \eta_{j+1}) \approx \phi(\eta_{j+1}; \eta_j + f_j \delta t, \Omega_j \delta t)$$

$$f_j = f(\eta_j, \tau_j), \Omega_j = (gg')( \eta_j, \tau_j)$$

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# Importance Sampling

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$$p(z_T, \dots, z_0) = \int p(z|\eta) \frac{p(\eta)}{p_2(\eta)} p_2(\eta) d\eta$$

- $p_2$ : importance density
- $p_{2,optimal} = \frac{p(z|\eta)p(\eta)}{p(z)} = p(\eta|z)$
- however,  $p(z)$  is unknown
- $\eta \rightarrow \eta(t, u)$ : random field,  $u$  = simulation time

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# Langevin Sampling

Langevin (1908); Roberts and Stramer (2002)

## Langevin equation

$$d\eta(u) = \partial_\eta \log p(\eta(u)|z)du + \sqrt{2}dW(u)$$

- stationary distribution of conditional latent states

$$p_{stat}(\eta) = e^{-\Phi(\eta)} = p(\eta|z)$$

- drift = -gradient of potential  $\Phi(\eta)$

$$-\partial_\eta \Phi(\eta) = \partial_\eta [\log p(z|\eta) + \log p(\eta) - \textcolor{red}{\log p(z)}]$$

sample from  $p(\eta|z)$ ,  $\textcolor{red}{p(z)}$  not needed!

- optimal nonlinear smoothing  
 $\eta(u) \sim p(\eta|z)$  in equilibrium  $u \rightarrow \infty$

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# Simulated Likelihood

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## Variance reduced MC-integration

$$\hat{p}(z_T, \dots, z_0) = L^{-1} \sum p(z|\eta_l) \frac{p(\eta_l)}{p_2(\eta_l)}$$

- $\eta_l \sim p(\eta|z)$  in equilibrium
- draw optimal  $\eta_l = \eta(u_l)$  from discretized Langevin equation (including Metropolis-Hastings mechanism)
- $p_{2,optimal} = \frac{p(z|\eta)p(\eta)}{p(z)} = p(\eta|z)$  is unknown
- Idea: estimate  $p_2 = p(\eta|z)$  from  $\eta_l \sim p(\eta|z)$

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# Estimation of importance density

- use known (suboptimal) reference density

$$p_2 = p_0(\eta|z) = p_0(z|\eta)p_0(\eta)/p_0(z)$$

- kernel density estimate

$$\hat{p}(\eta|z) = L^{-1} \sum_I k(\eta - \eta_I; \text{smooth})$$

Problem:

- high dimensional state,
- no structure imposed on  $p(\eta|z)$

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# Estimation of importance density

- use Markov structure of state space model

$$\eta_{j+1} = f(\eta_j)\delta t + g(\eta_j)\delta W_j$$

$$z_i = h(y_i) + \epsilon_i$$

$$p(\eta|z) = p(\eta_J|\eta_{J-1}, \dots, \eta_0, z) * p(\eta_{J-1}, \dots, \eta_0|z)$$

## Conditional Markov process

$$p(\eta_{j+1}|\eta_j, \dots, \eta_0, z) = p(\eta_{j+1}|\eta_j, z)$$

use conditional independence of **past**  $z^i = (z_0, \dots, z_i)$  and **future**  $\bar{z}^i = (z_{i+1}, \dots, z_T)$  given  $\eta^j$ :

$$p(\eta_{j+1}|\eta^j, z^i, \bar{z}^i) = p(\eta_{j+1}|\eta^j, \bar{z}^i) = p(\eta_{j+1}|\eta_j, \bar{z}^i)$$

$$p(\eta_{j+1}|\eta_j, z^i, \bar{z}^i) = p(\eta_{j+1}|\eta_j, \bar{z}^i)$$
$$j_i \leq j < j_{i+1}$$

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# Euler transition kernel

- Euler density (discretization interval  $\delta t$ )

$$p(\eta_{j+1}, \delta t | \eta_j) \approx \phi(\eta_{j+1}; \eta_j + f_j \delta t, \Omega_j \delta t)$$

- modified drift and diffusion matrix

conditional Euler density

$$\begin{aligned} p(\eta_{j+1}, \delta t | \eta_j, \textcolor{red}{z}) &\approx \\ \phi(\eta_{j+1}; \eta_j + (f_j + \delta f_j) \delta t, (\Omega_j + \delta \Omega_j) \delta t) \end{aligned}$$

- nonlinear regression for  $\delta f_j$  and  $\delta \Omega_j$   
(parametric and nonparametric)
- draw data  $\eta_{jl} = \eta(\tau_j, u_l) \sim p(\eta|z)$

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# Kernel density

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## conditional transition density

$$p(\eta_{j+1}, \delta t | \eta_j, z) = \frac{p(\eta_{j+1}, \eta_j | z)}{p(\eta_j | z)}$$

- estimate joint density  $p(\eta_{j+1}, \eta_j | z)$  and  $p(\eta_j | z)$  with **kernel density estimates**
- variant: use  $\phi(\eta_{j+1}, \eta_j | z)$  and  $\phi(\eta_j | z)$

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# Examples: Geometrical Brownian motion

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$$dy(t) = \mu y(t)dt + \sigma y(t) dW(t)$$

- nonlinear model (**multiplicative noise**)  $y * dW$
- exact solution: set  $x = \log y$ , use Itô's lemma

$$dx = dy/y + 1/2(-y^{-2})dy^2 = (\mu - \sigma^2/2)dt + \sigma dW$$

exact discrete model

$$y(t) = y(t_0)e^{(\mu - \sigma^2/2)(t-t_0) + \sigma[W(t) - W(t_0)]}$$

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Figure : Trajectory and log returns.

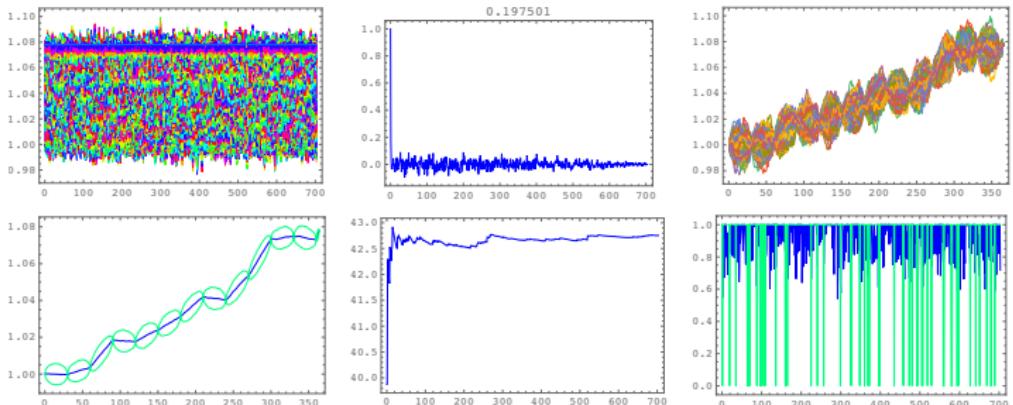


Figure : Langevin sampler,  $\hat{p}_2 = \prod_j \phi(\eta_{j+1}, \eta_j | z) / \phi(\eta_j | z)$ .

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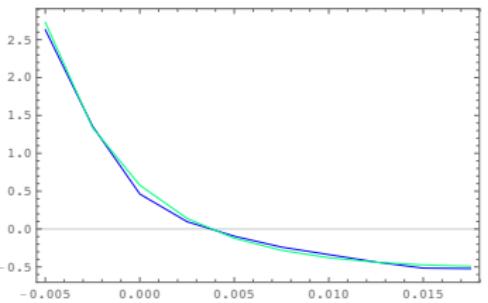
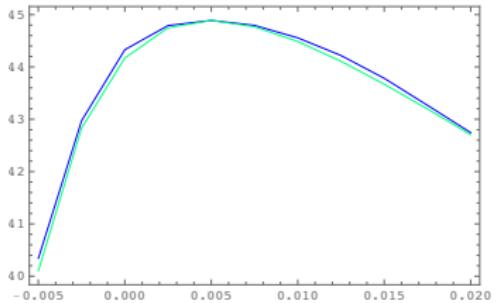


Figure : likelihood and score,  $\hat{p}_2$  = conditional kernel density.

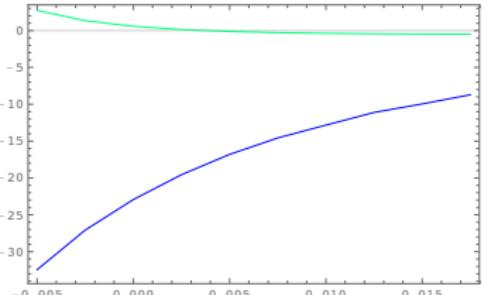
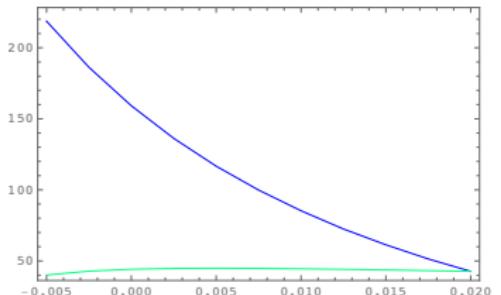


Figure : likelihood and score,  $\hat{p}_2$  = full kernel density.

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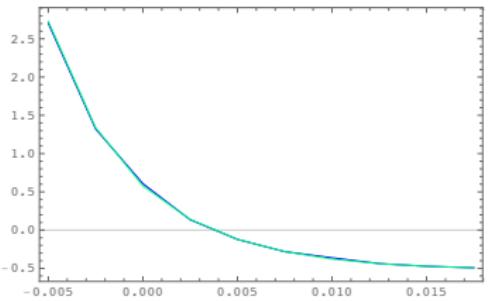
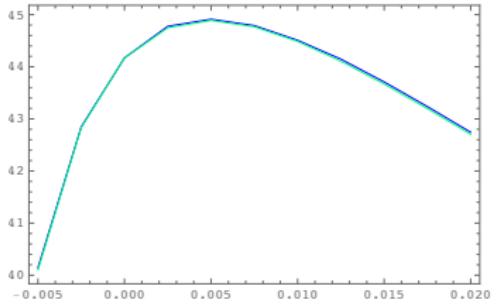


Figure : likelihood and score,  $\hat{p}_2$  = conditionally Gaussian.

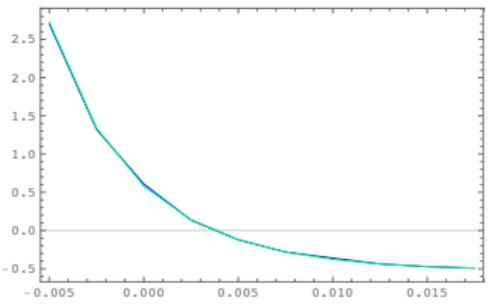
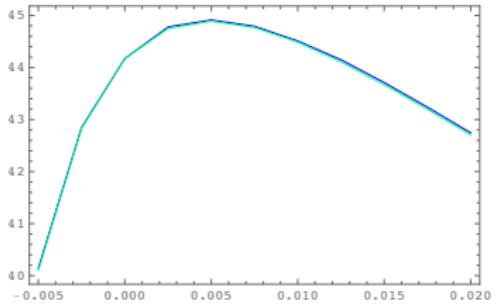


Figure : likelihood and score,  $\hat{p}_2$  = linear GLS.

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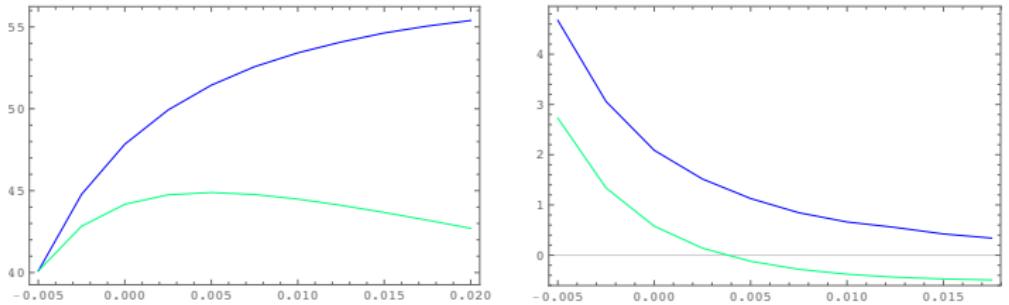


Figure : likelihood and score,  $\hat{p}_2 = \text{linear GLS}$ ,  
constant diffusion matrix.

# Examples: Cameron-Martin formula

$$E\left[e^{-\frac{\lambda^2}{2} \int_0^T W(t)^2 dt}\right] = 1/\sqrt{\cosh(T\lambda)}$$

Cameron and Martin (1945); Gelfand and Yaglom (1960)

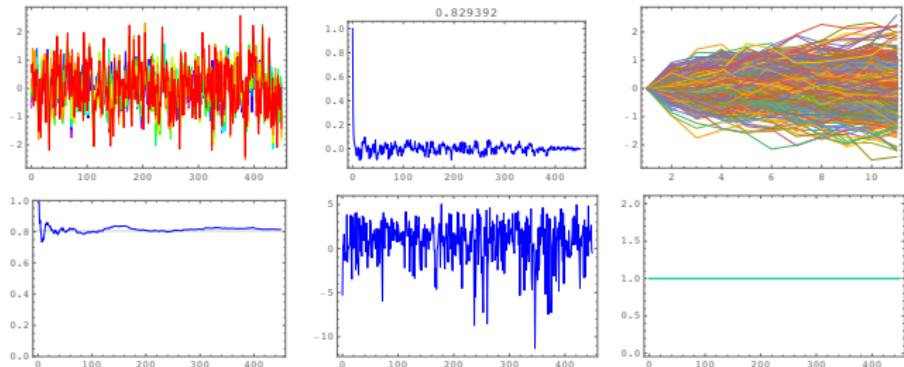


Figure : Cameron-Martin formula.

Simulation using a conditionally gaussian importance density.  
 $T = 1, \lambda = 1, dt = 0.1$  and  $L = 500$  replications. Exact value  
 $1/\sqrt{\cosh(1)} = 0.805018$

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# Examples: Cameron-Martin formula

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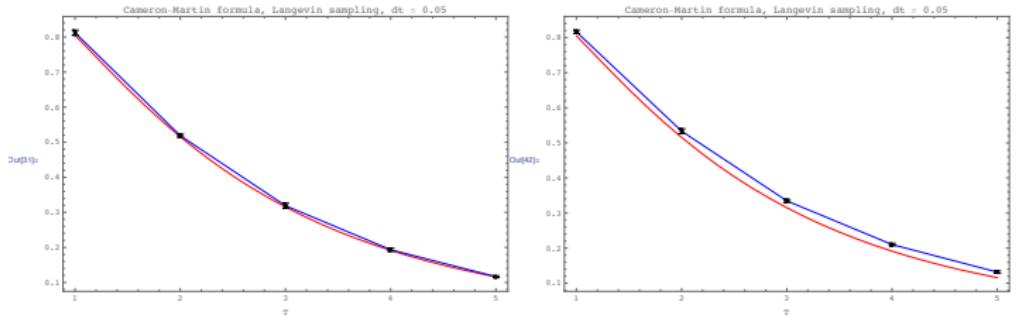


Figure : Expectation value as a function of  $T$ . Right:  $\Omega = \text{fix.}$

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## Examples: Feynman-Kac-formula

### Schrödinger Equation (imaginary time $t = i\tau$ )

$$u_t = \frac{1}{2}u_{xx} - \phi(x)u$$

$$\begin{aligned} u(x, t=0) &= \delta(x-z) \\ \phi(x) &= \frac{1}{2}\gamma^2x^2 \end{aligned}$$

### Feynman-Kac-formula

$$\begin{aligned} u(x, t) &= E_x \left[ e^{-\frac{\gamma^2}{2} \int_0^t W(u)^2 du} \delta(W(t) - z) \right] = \\ &\sqrt{\frac{\gamma}{2\pi \sinh(\gamma t)}} \times \\ &\exp \left( \frac{\gamma}{2 \sinh(\gamma t)} [2xz - (x^2 + z^2) \cosh(\gamma t)] \right) \end{aligned}$$

# Feynman-Kac-formula

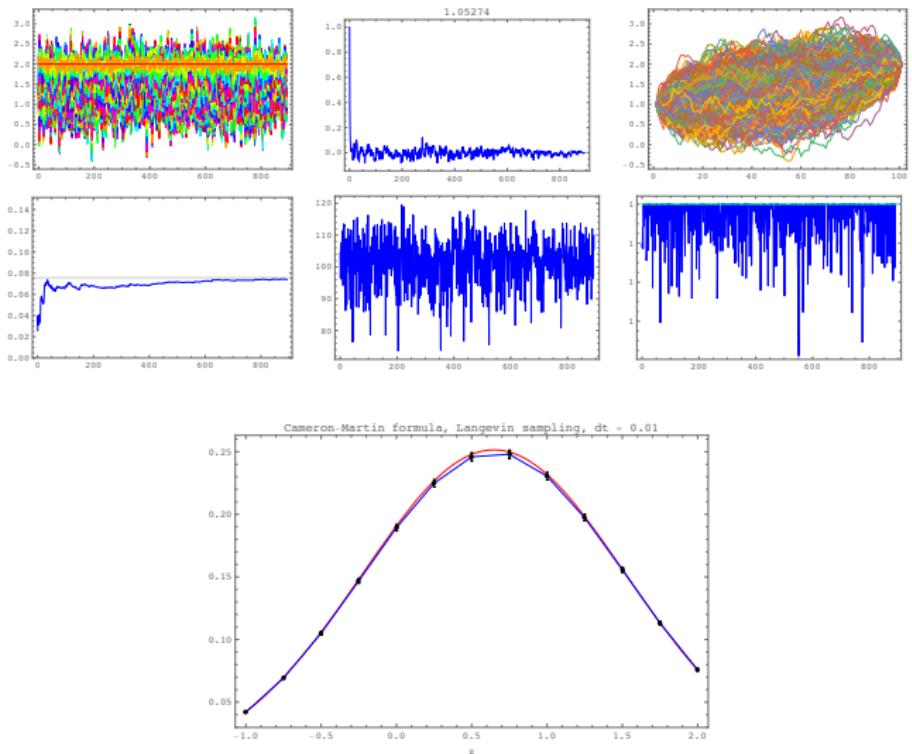


Figure : Expectation value as a function of  $z$ ,  $x = 1$ ,  $t = 1$ ,  $\gamma = 1$ .

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# Conclusion

- **smooth likelihood simulation** for nonlinear continuous-discrete state space models.
- **Nonlinear smoothing** of latent variables between measurements.
- **Variance reduced MC estimation of functional integrals** in finance, statistics and quantum theory (Feynman-Kac formula).

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